

FITTING A GAUSSIAN MIXTURE MODEL THROUGH THE GINI INDEX

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A linear combination of Gaussian components is known as a Gaussian mixture model. It is widely used in data mining and pattern recognition. In this paper, we propose a method to estimate the parameters of the density function given by a Gaussian mixture model. Our proposal is based on the Gini index, a methodology to measure the inequality degree between two probability distributions, and consists in minimizing the Gini index between an empirical distribution for the data and a Gaussian mixture model. We will show several simulated examples and real data examples, observing some of the properties of the proposed method.

Keywords: Gini index problem, Gaussian mixture model, clustering.

1. Introduction

Consider the problem of finding clusters for the data set $P = \{p_1, p_2, \dots, p_M\}$ with $p_m \in \mathbb{R}^N = X$. When we want to analyze the data set by modelling their behaviour, we usually use some of the known density functions, for example the multivariate normal density of dimension N given by

$$f(x|\mu, \Sigma) = \frac{1}{(2\pi)^{N/2} |\Sigma|^{1/2}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1} (x-\mu)}, \quad (1)$$

where μ is the mean and Σ is the covariance matrix. However, sometimes a unimodal distribution cannot represent the information given by the data when there are clusters, so multimodal distributions are used.

In this paper we consider a multimodal distribution that is a linear combination of Gaussian components to model the data, known as a Gaussian mixture model, (Bishop, 2006; Reynolds, 2009). The Gaussian mixture model is widely used for segmentation of images, speech recognition, language identification and statistical representation (Greenspan *et al.*, 2006; Povey *et al.*, 2011; Torres-Carrasquillo *et al.*, 2002; Singh *et al.*, 2009).

In the Gaussian mixture model we consider a density function that is a linear combination of K Gaussian

densities of the form

$$\sum_{k=1}^K \phi_k f(x|\mu_k, \Sigma_k).$$

The components of this mixture are Gaussian densities $f(x | \mu_k, \Sigma_k)$ with mean μ_k and covariance matrix Σ_k , and the ϕ_k parameters are mixing coefficients that must comply with

$$0 \leq \phi_k \leq 1, \quad k = 1, \dots, K, \quad (2)$$

and

$$\sum_{k=1}^K \phi_k = 1. \quad (3)$$

The sought parameters in the Gaussian mixture distribution are $\phi = (\phi_1, \phi_2, \dots, \phi_K)$, $\mu = (\mu_1, \mu_2, \dots, \mu_K)$ and $\Sigma = (\Sigma_1, \Sigma_2, \dots, \Sigma_K)$. One way to set the values of these parameters is to use the maximum likelihood method. The logarithm of the likelihood function for this problem, when we have M elements, that is, $\{p_m\}_{m=1}^M$, is given by

$$l(\phi, \mu, \Sigma|P) = \sum_{m=1}^M \log \left(\sum_{k=1}^K \phi_k f(p_m | \mu_k, \Sigma_k) \right).$$

There is no analytical solution to this problem, so iterative numerical optimization techniques are used for

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this purpose. In several texts (e.g., Dempster *et al.*, 1977; Meng and Rubin, 1994; Vaida, 2005; Xu and Jordan, 1996), the authors employ a powerful framework called expectation-maximization for Gaussian mixtures. They want to maximize the likelihood function with respect to the parameters ϕ , μ and Σ by following the EM algorithm.

Once the desired ϕ , μ and Σ parameters are obtained, we can perform data classification, using the total probability law and the Bayes theorem. We can obtain the probability that, given a data point x , it belongs to the parametric distribution g_k , that is, $\Pr(g_k|x)$, for $k = 1, \dots, K$, in the following way:

$$\Pr(g_{k'}|x) = \frac{\Pr(g_{k'})\Pr(x|g_{k'})}{\sum_{k=1}^K \Pr(g_k)\Pr(x|g_k)},$$

for $k' = 1, 2, \dots, K$, where $\Pr(x|g_k)$ is the probability that x comes from the parametric distribution k and $\Pr(g_k) = \phi_k$ is the probability of the parametric distribution k . Once we obtain the probabilities $\Pr(g_k|x)$, with $k = 1, \dots, K$, we make a comparison of these values and we determine that the point x belongs to the parametric distribution with the greatest value $\Pr(g_k|x)$. In this work we propose an alternative to the EM algorithm, called the GI algorithm, based on minimizing the Gini index between the empirical distribution of the observed data and a mixture of Gaussians.

In Section 2, we will give a brief introduction to the Gini index. In Section 3, we will show the proposed procedure to estimate the parameters of a Gaussian mixture model through the Gini index problem in an efficient way, similarly to but in greater detail than in our previous work (López-Lobato and Avendaño-Garrido, 2020). In Section 4 we will perform exhaustive experiments with simulated data and a few real data sets, with the purpose of comparing the numerical results obtained by the EM algorithm, the K -means method and the algorithm proposed in this work. We will end by giving conclusions and mentioning future work in Section 5.

2. Gini index

The Gini index is a distance between two probability distributions, so it can be used to measure the inequality level between them. This measure is applied as an indicator of social and economic inequality when the income distribution of a country is analyzed and in other several fields like engineering, transport and ecology (see Giorgi and Gigliarano, 2017; Rachev *et al.*, 2013; Ultsch and Lötsch, 2017).

For the Gini index (GI) problem, we consider a space X , two probability distributions ν_1 and ν_2 on X , and a distance function in X , $d : X \times X \rightarrow \mathbb{R}$.

The GI problem is as follows:

Minimize

$$\int_{X \times X} d(x, y) d\pi$$

subject to

$$\begin{aligned} \pi(A \times X) &= \nu_1(A), \\ \pi(X \times A) &= \nu_2(A), \\ \pi &\in \mathcal{M}^+(X \times X), \end{aligned}$$

for all A in the Borel σ -algebra of X , where $\mathcal{M}^+(X \times X)$ is the convex cone of probability measures, i.e., π is in the set of joint probability distributions on $X \times X$, whose marginals in the first and second components are the probability distributions ν_1 and ν_2 , respectively, denoted by $\Pi(\nu_1, \nu_2)$.

Rubner *et al.* (2000) and Villani (2003) showed that this problem always has a solution, which is a distance between the probability distributions ν_1 and ν_2 and can be very expensive to find. In addition, the solution $\pi^* \in \Pi(\nu_1, \nu_2)$ is a probability measure and the optimal value of this problem defines the Gini index between the probability measures ν_1 and ν_2 , denoted by $GI(\nu_1, \nu_2)$, that is,

$$GI(\nu_1, \nu_2) = \int_{X \times X} d(x, y) d\pi^*.$$

In this paper, we propose to find the Gaussian mixture distribution that minimizes the Gini index to an empirical distribution. We make this proposal based on the work of Bassetti *et al.* (2006), where the theory of the minimum dissimilarity estimators and the estimators of the minimum distance of Kantorovich are discussed. In this type of problems, the distribution ν_1 is known, commonly associated with an empirical distribution, and the distribution ν_2 must be estimated in such a way that the distance between ν_1 and ν_2 , given by the Gini index, is minimum. We propose this in order to efficiently estimate the parameters of a Gaussian mixture model through the Gini index problem.

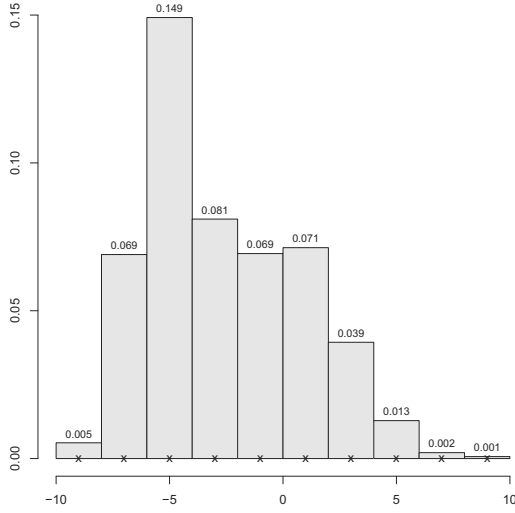
Now, we will explain the process that we follow to establish the Gini Index problem based on a given data set and the way in which we solved the problem.

3. Parameter estimation minimizing the Gini index

3.1. Empirical distribution ν_1 . We suppose to have a data set

$$P = \left\{ p_m = \left(p_1^{(m)}, p_2^{(m)}, \dots, p_N^{(m)} \right) \right\}_{m=1}^M,$$

with M elements in dimension N . Consider this set P as the following data frame:

Fig. 1. Representative histogram for column C_n .

	C_1	C_2	$\dots$	C_N
$p_1 =$	$p_1^{(1)}$	$p_2^{(1)}$	$\dots$	$p_N^{(1)}$
$p_2 =$	$p_1^{(2)}$	$p_2^{(2)}$	$\dots$	$p_N^{(2)}$
$\vdots$	$\vdots$	$\vdots$	$\ddots$	$\vdots$
$p_M =$	$p_1^{(M)}$	$p_2^{(M)}$	$\dots$	$p_N^{(M)}$

For each $n = 1, \dots, N$, we define data frame column sets as

$$C_n = \left\{ p_n^{(m)} \right\}_{m=1}^M,$$

and we obtain a representative histogram of each of them. We assume that $C_n \in [\alpha_n, \beta_n]$. The histogram helps us to divide the set $[\alpha_n, \beta_n]$ into several bins. We use the count of these bins as a density estimate. If we have S_n bins, the partition is

$$B_1^n = [y_0^{(n)}, y_1^{(n)}), \quad B_2^n = [y_1^{(n)}, y_2^{(n)}), \\ \dots, \quad B_{S_n}^n = [y_{S_n-1}^{(n)}, y_{S_n}^{(n)}].$$

The variable $Y_n = \left\{ y_j^{(n)} \right\}_{j=0}^{S_n}$ is represented by the x 's in Fig. 1. In such a case, for a given point $x_n \in B_j^n$ we define the density estimation as

$$\hat{\gamma}_n(x_n) = \frac{1}{M} \sum_{m=1}^M I_{B_j^n}(p_n^{(m)}),$$

where I_A is the indicator function of the set A .

To define the empirical distribution ν_1 in X , we consider the multiplication of the density estimation for each of the columns, i.e., for $x = (x_1, x_2, \dots, x_N) \in X$ we have

$$f_1(x) = \prod_{n=1}^N \hat{\gamma}_n(x_n). \quad (4)$$

3.2. Parametric distribution ν_2 . For the parametric distribution ν_2 on X , we consider a Gaussian mixture density, that is,

$$f_2(x) = \sum_{k=1}^K \phi_k f(x|\mu_k, \Sigma_k).$$

The parameters ϕ_k are the mixture proportions and must comply with (2) and (3). In this work, the function f is assumed that is an *independent multivariate normal distribution* of dimension N , given in (1), where the mean μ_k is a real vector $[\mu_{k1}, \mu_{k2}, \dots, \mu_{kN}]^T$ and the covariance matrix Σ_k is a real diagonal positive definite $N \times N$, matrix i.e.,

$$\Sigma_k = \text{diag}(\sigma_{k1}^2, \sigma_{k2}^2, \dots, \sigma_{kN}^2). \quad (5)$$

Then, by the form of Σ_k , we know that

$$f(x|\mu_k, \Sigma_k) = \prod_{n=1}^N g(x_n|\mu_{kn}, \sigma_{kn}^2),$$

for $x = (x_1, x_2, \dots, x_N) \in X$, where g is the univariate normal density function, that is,

$$g(s|\mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(s-\mu)^2}{2\sigma^2}}.$$

Thus, we have, for $x = (x_1, \dots, x_N) \in X$,

$$f_2(x) = \sum_{k=1}^K \phi_k \cdot \prod_{n=1}^N g(x_n|\mu_{kn}, \sigma_{kn}^2). \quad (6)$$

Remark 1. It should be noted that we use independent probability densities in order not to increase the computational cost of the proposed model. This fundamental assumption is made to simplify the real-world problem complexity in a similar way to the naive Bayesian classification model (Mao *et al.*, 2020), and take individual problems for each coordinate in the multidimensional case as done by Kulczycki (2018). In several works (e.g., Elkan, 1997; Flach and Lachiche, 2004), it has been mentioned that even with this unrealistic assumption this technique is effective in practice.

3.3. Gini index problem for a Gaussian mixture model. The Gini index problem is as follows:

Minimize

$$\int_{X \times X} d(x, y) d\pi$$

subject to

$$\begin{aligned}\pi(A, X) &= \int_A f_1(x) dx, \\ \pi(X, A) &= \int_A f_2(x) dx \\ \pi &\in \mathcal{M}^+(X \times X),\end{aligned}$$

for all A in the Borel σ -algebra of X , where $\mathcal{M}^+(X \times X)$ in the set of joint probability distributions on $X \times X$, d is the Euclidean distance in X , f_1 is as in (4) and f_2 is as in (6). We are looking for the following parameters of the density f_2 given in (6):

- the proportions defining the mixture $\phi = (\phi_1, \phi_2, \dots, \phi_K)$, which comply with (2) and (3);
- the means of the Gaussian components $\mu = (\mu_1, \mu_2, \dots, \mu_K)$;
- the covariance matrices of the independent Gaussian components $\Sigma = (\Sigma_1, \Sigma_2, \dots, \Sigma_K)$ as in (5).

3.4. Parameter estimation. As the solution exists, by Bassetti *et al.* (2006), we can solve this problem using the Lagrange multiplier method. The Lagrangian for this problem is

$$\begin{aligned}\mathcal{L}(\phi, \mu, \Sigma) &= \sum_{x \in X} \sum_{y \in X} d(x, y) \pi(x, y) \\ &\quad - \sum_{x \in X} \lambda_x \left[\sum_{y \in X} \pi(x, y) - f_1(x) \right] \\ &\quad - \sum_{y \in X} \gamma_y \left[\sum_{x \in X} \pi(x, y) - f_2(y) \right] \\ &\quad - \alpha \left[\sum_{x \in X} \sum_{y \in X} \pi(x, y) - 1 \right] \\ &\quad - \beta \left[\sum_{k=1}^K \phi_k - 1 \right].\end{aligned}\quad (7)$$

Consider first setting the values for the joint distribution π . We write $\pi_{xy} = \pi(x, y)$ and $d_{xy} = d(x, y)$. For s and t fixed at X , the following is true:

$$\frac{\partial \mathcal{L}}{\partial \pi_{st}} = d_{st} - \lambda_s - \gamma_t - \alpha = 0.$$

Then it is true that $d_{st} = \lambda_s + \gamma_t + \alpha$. Since d is a distance in X , if $s = t$, we have that $\gamma_t = -\lambda_t - \alpha$, due to $d_{tt} = 0$. Also, as $d_{ts} = d_{st}$, it is true that $\lambda_t + \gamma_s + \alpha = \lambda_s + \gamma_t + \alpha$, and then $\lambda_t - \lambda_s = -(\lambda_t - \lambda_s)$. This equality is fulfilled if and only if $\lambda_t = \lambda_s$, for $s, t \in X$.

Similarly, it is true that $\gamma_t = \gamma_s$ for $s, t \in X$. Then $\lambda_x = \lambda$ for every $x \in X$, $\gamma_y = \gamma$ for every $y \in X$ and $\gamma = -\lambda - \alpha$. With these equalities, the Lagrangian (7) turns out to be

$$\begin{aligned}\mathcal{L}(\phi, \mu, \Sigma) &= \sum_{x \in X} \sum_{y \in X} d_{xy} \pi_{xy} + \lambda \sum_{x \in X} f_1(x) \\ &\quad - (\lambda + \alpha) \sum_{y \in X} f_2(y) + \alpha - \beta \left[\sum_{k=1}^K \phi_k - 1 \right].\end{aligned}\quad (8)$$

To obtain the means, we differentiate the Lagrangian (8) with respect to μ_{tr} , with fixed $1 \leq t \leq K$ and fixed $1 \leq r \leq N$. Because the only sum in which the variable μ_{tr} appears is where ν_2 is present, we have

$$\frac{\partial \mathcal{L}}{\partial \mu_{tr}} = -(\lambda + \alpha) \sum_{y \in X} \frac{\partial}{\partial \mu_{tr}} \nu_2(y)$$

In order to use the definition of f_2 given in (6), we must set the sum of all the elements of X given in the above equation in terms of their coordinates, so

$$\begin{aligned}& -(\lambda + \alpha) \sum_{y \in X} \frac{\partial}{\partial \mu_{tr}} f_2(y) \\ &= -(\lambda + \alpha) \sum_{j_1=1}^{S_1} \sum_{j_2=1}^{S_2} \dots \sum_{j_N=1}^{S_N} \frac{\partial}{\partial \mu_{tr}} \left[\sum_{k=1}^K \phi_k \cdot \prod_{n=1}^N g(y_{j_n}^{(n)} | \mu_{kn}, \sigma_{kn}^2) \right].\end{aligned}$$

Thus, we have

$$\begin{aligned}\frac{\partial \mathcal{L}}{\partial \mu_{tr}} &= -(\lambda + \alpha) \cdot \phi_t \left(\frac{1}{\sigma_{tr}^2} \right) \\ &\quad \cdot \sum_{j_1=1}^{S_1} \sum_{j_2=1}^{S_2} \dots \sum_{j_N=1}^{S_N} \left[\prod_{n=1}^N g(y_{j_n}^{(n)} | \mu_{tn}, \sigma_{tn}^2) \right. \\ &\quad \cdot \left. (y_{j_r}^{(r)} - \mu_{tr}) \right] = 0.\end{aligned}$$

Here $\lambda + \alpha$ must be different from 0, since if it were 0, γ would be 0 and there would be no restrictions. In the same way, ϕ_t must be different from 0, since if it were, there would be no K Gaussian components. Accordingly,

$$\begin{aligned}& \sum_{j_1=1}^{S_1} \sum_{j_2=1}^{S_2} \dots \sum_{j_N=1}^{S_N} \left[\prod_{n=1}^N g(y_{j_n}^{(n)} | \mu_{tn}, \sigma_{tn}^2) \right. \\ &\quad \cdot \left. (y_{j_r}^{(r)} - \mu_{tr}) \right] = 0.\end{aligned}$$

Then we have

$$\begin{aligned}\mu_{tr} &= \frac{\sum_{j_1=1}^{S_1} \sum_{j_2=1}^{S_2} \dots \sum_{j_N=1}^{S_N} y_{j_r}^{(r)} \prod_{n=1}^N g(y_{j_n}^{(n)} | \mu_{tn}, \sigma_{tn}^2)}{\sum_{j_1=1}^{S_1} \sum_{j_2=1}^{S_2} \dots \sum_{j_N=1}^{S_N} \prod_{n=1}^N g(y_{j_n}^{(n)} | \mu_{tn}, \sigma_{tn}^2)}.\end{aligned}$$

Simplifying the common terms in the numerator and denominator, we get

$$\mu_{tr} = \frac{\sum_{j_r=1}^{S_r} y_{j_r}^{(r)} \exp\left(-\frac{(y_{j_r}^{(r)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}{\sum_{j_r=1}^{S_r} \exp\left(-\frac{(y_{j_r}^{(r)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}. \quad (9)$$

If we differentiate (8) with respect to σ_{tr} , with fixed $1 \leq t \leq K$, $1 \leq r \leq N$, analogously to the previous case, we get

$$\sigma_{tr}^2 = \frac{\sum_{j_r=1}^{S_r} (y_{j_r}^{(r)} - \mu_{tr})^2 \exp\left(-\frac{(y_{j_r}^{(r)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}{\sum_{j_r=1}^{S_r} \exp\left(-\frac{(y_{j_r}^{(r)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}. \quad (10)$$

The expressions (9) and (10) can be evaluated iteratively. This estimate is made with respect to the data set P , so the iterative expressions of the *GI algorithm* are

$$\mu_{tr} = \frac{\sum_{m=1}^M p_r^{(m)} \exp\left(-\frac{(p_r^{(m)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}{\sum_{m=1}^M \exp\left(-\frac{(p_r^{(m)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}, \quad (11)$$

for $1 \leq t \leq K$ and $1 \leq r \leq N$,

$$\sigma_{tr}^2 = \frac{\sum_{m=1}^M (p_r^{(m)} - \mu_{tr})^2 \exp\left(-\frac{(p_r^{(m)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}{\sum_{m=1}^M \exp\left(-\frac{(p_r^{(m)} - \mu_{tr})^2}{2\sigma_{tr}^2}\right)}, \quad (12)$$

for $1 \leq t \leq K$ and $1 \leq r \leq N$.

For the proportions of the mixture ϕ_k , as in the EM algorithm, we use in each iteration the expression (1) to calculate for each point p_m , $m = 1, \dots, M$, the probability that p_m comes from the parametric distribution k , i.e., $P(f_k|p_m)$, with $k = 1, \dots, K$, where $f_k = f(x|\mu_k, \Sigma_k)$. Then we can compare these values and determine the membership of the point p_m . Once this classification is made, we can obtain the proportions by taking the quantity of elements in each class and normalizing via dividing by M .

It is important to emphasize that, to obtain the expressions (11) and (12), we made the assumption that the components of the Gaussian mixture f_2 were independent multivariate normal distributions, and for this reason we only looked for the values of the covariance matrix that is a positive definite diagonal matrix. In the following section we will show some of the results obtained when using this algorithm.

4. Numerical results

In this section we will show the obtained numerical results when using the EM algorithm, the GI algorithm and the K -means method. We consider experiments with simulated data and real data. In the experiments with simulated data we consider data from 2, 3 and 4 Gaussians. For experiments with real data we consider two databases: the Iris data set and the Seeds data set, found in the UCI Machine Learning Repository¹.

To carry out these experiments we used the free software R². For the classification with the K -means method we use the `kmeans()` function, and for the classification using the EM algorithm we employ the `GMM()` function in the `ClusterR` package.

With respect to the initial values that we use for the GI algorithm, we select them as follows. For the initial values of the means, we make a random selection of points within the domain of the analysed data set. For the initial values of the covariance matrix, we randomly select values between 1 and 4. For the initial values of the mixing proportions, we use, as usual, the values of a uniform distribution, that is, if we want adjust K Gaussians, we take proportions $\phi_k = 1/K$, for each k , as the initial values.

4.1. Simulated data. First we will show how the training data was generated and then what process will be used to make comparisons between the GI algorithm, the EM algorithm and the K -means model.

We generated a training data set P with a sample size $M = 3000$, from 2, 3 or 4 classes when considering different configurations for each case (*univariate* and *bivariate*), by analyzing the following characteristics:

- data proportion: the data amount from each class can be *equal* or *different*;
- data intersection: classes can be *spatially well differentiated* or *intersected*.

We specify the values used in each configuration in the corresponding sections. Once the data set to be analyzed has been generated, we translate the data items to the interval $[-10, 10]$ for the univariate case, and to $[-10, 10] \times [-10, 10]$ for the bivariate case. In addition, we can obtain a class vector that establishes from which Gaussian each data item comes from, in order to be able to establish the classification accuracy with the analysed models.

For the experiments with synthetic data we perform the following process 100 times:

1. We generate a set P of 3000 training data items specifying the characteristics of the configurations.

¹<https://archive.ics.uci.edu/ml/index.php>

²<https://www.r-project.org/>.

Table 1. Results of Configuration 1: univariate case.

Adjusted Gaussians	Acc.	GI algorithm	EM algorithm	K-means
2	Acc. time GI	100 0.58 0.300322	100 0.0345	99.96667 0.0328
3	Acc. time GI	100 0.5683 0.300322	78.8483333 1.3175	75.66667 0.2292
4	Acc. time GI	100 0.6651 0.300322	77.0606667 1.643	74.3270 0.3069

2. We adjust each of the three models to the data set P , assuming that there are 2, 3, 4 or 5 groups, depending on each case. Then we obtain the classification accuracy.
3. We record the obtained classification accuracy through the three models and, for the univariate case, the Gini index values given by the GI algorithm.

Once we obtain the results of 100 iterations of these experiments, we calculate the average classification accuracy for each model; the time average, and the average values of the Gini index in the univariate case. These values are recorded in the tables in the following sections. In each table in this document, we show in boldface the best classification accuracy average. Let us recall that a small value for the Gini index suggests that the empirical distribution is closer to the estimated theoretical model.

4.1.1. Univariate case.

Configuration 1. (See Fig. 2(a)) We have 3000 data items generated with *equal proportions* of 2 *separate univariate* Gaussians, with the following parameters: 1500 data items with $\mu_1 = -12$, $\sigma_1 = 2$ and 1500 data items with $\mu_2 = 10$, $\sigma_2 = 3$. In this case we adjust 2, 3 and 4 Gaussians with the three methods, and the results are shown in Table 1.

The best classification accuracy average for the three models are those that consider the search for two Gaussians, because the data was generated precisely from two Gaussians. With the GI algorithm and the EM algorithm we obtain a 100% of average accuracy, higher than the K -means method. We obtain 100% when searching for 3 and 4 Gaussians with the GI algorithm because, even when it finds them, one or two have a proportion $\phi_k = 0$. For the value of the Gini index, we can see that the minimum is obtained for 2 Gaussians, which is an advantage, since we do not require *a priori* knowledge about the number of components in the mixture.

Table 2. Results of Configuration 2: univariate case.

Adjusted Gaussians		GI algorithm	EM algorithm	K-means
3	Acc. time GI	87.5 0.5689 1.408632	86.732 0.0995	81.0850 0.2305
4	Acc. time GI	99.899333 0.4542 1.109472	94.36667 1.6996	99.86667 0.3029
5	Acc. time GI	99.899333 0.6337 1.109472	83.776333 2.0762	89.8310 0.3678

Configuration 2. (See Fig. 2(b)) We have 3000 data points generated with *different proportions* of 4 *intersected univariate* Gaussians, with the following parameters: 375 data points with $\mu_1 = -22$, $\sigma_1 = 3$, 750 data points with $\mu_2 = -8$, $\sigma_2 = 2$, 750 data points with $\mu_3 = 5$, $\sigma_3 = 2$ and 1125 data points with $\mu_4 = 15$, $\sigma_4 = 1$. In this case we adjust 3, 4 and 5 Gaussians. In Table 2 we show the results for this configuration.

In this configuration the GI algorithm yielded the best result when considering 4 and 5 Gaussians, and the value of the Gini index is lower in these cases. Again, we find a proportion $\phi_k = 0$ when we adjust 5 Gaussians with the GI algorithm.

Configuration 3. (See Fig. 2(c)) We have 3000 data items generated with *equal proportions* of 3 *univariate* distributions, 2 *intersected*, with the following characteristics: 1000 data items from a beta distribution $X \sim \beta(2, 1.5)$, 1000 data items from a chi-square distribution $X \sim \chi^2(40)$ and 1000 data items from a Poisson distribution $X \sim \text{Pois}(100)$. We adjust 2, 3 and 4 Gaussians in this case. The results are shown in Table 3.

In this configuration we obtained a better classification when we adjusted 3 and 4 Gaussians with the GI algorithm. With this configuration we can see that, even when we generate the data set to be analyzed from non-Gaussian distributions, the method works correctly, that is, it is robust in this regard.

Configuration 4. (See Fig. 2(d)) We have 3000 data points generated with *different proportions* from 3 *intersected univariate* distributions, with the following characteristics: 1500 data points from a gamma distribution $X \sim \Gamma(0.7, 0.3)$, 1000 data points from a binomial distribution $X \sim \text{Bin}(20, 0.8)$, 500 data points from a Poisson distribution $X \sim \text{Pois}(35)$. The obtained results for this configuration are found in Table 4.

In this case, we obtained a better classification with the K -means method.

In Configurations 3 and 4 we obtain the same averages for 4 Gaussians with the GI algorithm, which

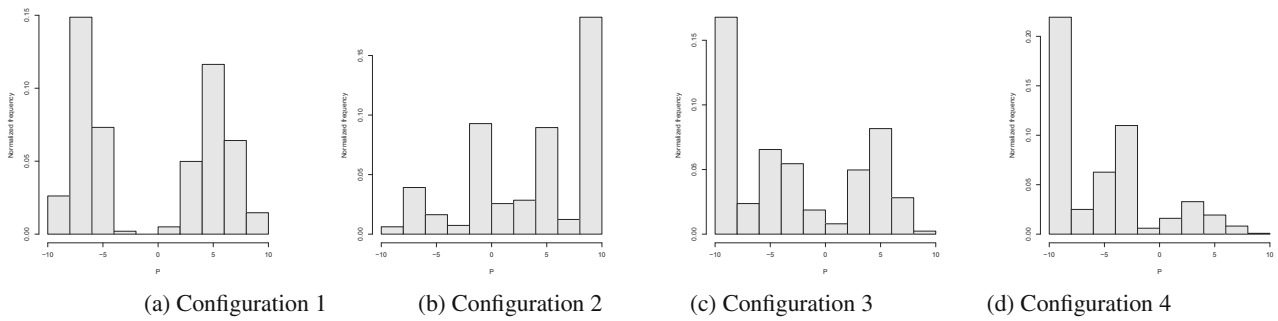


Fig. 2. Configuration examples for the univariate case.

Table 3. Results of Configuration 3: univariate case.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc.	66.666667	66.633	66.66667
	time	2.1092	0.1714	0.129
	GI	3.810997		
3	Acc.	99.497333	98.53267	94.7140
	time	3.3735	0.1525	0.8571
	GI	3.32223		
4	Acc.	99.497333	83.836333	76.66667
	time	3.5821	5.3871	1.1383
	GI	3.32223		

Table 4. Results of Configuration 4: univariate case.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc.	79.659	75.966667	81.63367
	time	0.3753	0.0779	0.033
	GI	1.987989		
3	Acc.	97.901333	97.501	98.03333
	time	1.0017	0.1584	0.2376
	GI	1.401458		
4	Acc.	97.901333	88.364667	87.26267
	time	1.055	0.4704	0.3044
	GI	1.401458		

means that in both cases the model found the same means and deviations values for 3 and 4 Gaussians, with a mixing proportion $\phi_k = 0$ for the case of 4 Gaussians. Although the results using the GI algorithm are not the best for Configuration 4, they are very close to those obtained with the K -means method, with the additional information of the covariance matrices.

These last two configurations have the ability to verify how the algorithm behaves with other distributions, not only for Gaussian mixtures. As can be seen, we obtain similar results to experiments with data generated from Gaussian mixtures.

In order to complement the experiments carried out in this section, we will analyze other situations that could be of interest, considering experiments with different amounts of data and different numbers of intervals in the representative histogram.

Different amount of data. In Tables 5 and 6, we show the results that we obtain by using the same parameters of Configurations 2 and 3, respectively, but varying the amount of data. In each table we show the data set size.

It is important to mention that the proportions used for the generation of the data correspond to the configuration proportions, i.e., the proportions 1/8, 1/4, 1/4 and 3/8 for Configuration 2 and a uniform distribution for Configuration 3.

As you can see in Table 5 in addition to Table 2, the

results given by the GI algorithm are not affected by the quantity of data. The values for the Gini index change in each case, but the minimum value always appears when we adjust 4 Gaussians. In Tables 3 and 6 the same thing is noticeable as in the previous case. It is important to note that, in this case, when we have less data, i.e., 150 elements, we obtain the best results. Again, we have the same accuracy for 3 and 4 Gaussians and the lowest Gini index in the same cases, because in the case of 4 Gaussians one of them has a zero proportion.

Different numbers of intervals in the representative histogram. For the experiments with different numbers of intervals in the representative histogram, we consider Configuration 2, with 4, 20 and 40 intervals. The results for this experiments are in Table 7. In Table 2, we show the results for 10 intervals.

As can be seen, the results do not depend on the number of intervals. The only value that varies with respect to the number of intervals is the Gini index. The lowest value appears when we adjust 4 Gaussians.

In all the experiments carried out in this work, we use the histogram generated by default by the `hist()` function of R, with 10 intervals.

4.1.2. Bivariate case. For the data generation in the plane, we consider independent bivariate normal distributions, that is, Gaussians whose covariance matrix

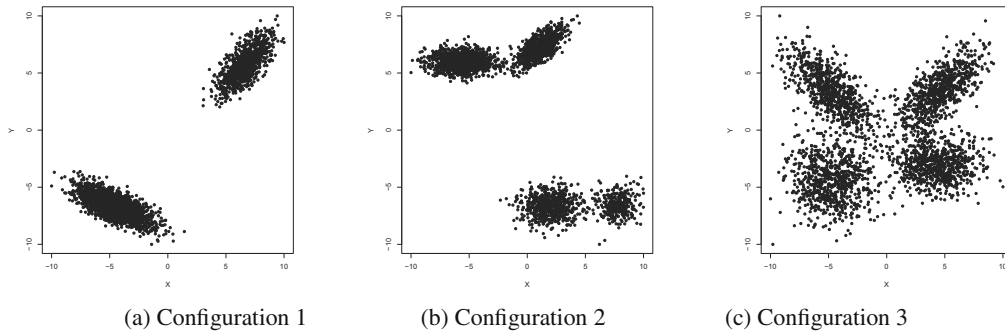


Fig. 3. Configuration examples for the bivariate case.

Table 5. Results of different amount of data for Configuration 2.

Set size	Adj. G		GI algorithm	EM algorithm	K-means
160	3	Ac. time GI	86.875 0.2451 1.561819	86.5 0.0227	86.875 0.05
	4	Acc. time GI	99.875 0.0794 1.184991	98.75 0.0527	99.3750 0.0674
	5	Acc. time GI	99.375 0.1978 1.197076	88.125 0.0715	84.3750 0.0748
320	3	Ac. time GI	87.5 0.4076 1.494086	85 0.0412	86.86875 0.14730
	4	Acc. time GI	99.75 0.254 1.158949	99.125 0.05430	99.41875 0.1979
	5	Acc. time GI	99.75 0.2487 1.15895	86.790625 0.704	84.10938 0.2149
1280	3	Ac. time GI	85.12344 1.45510 1.41227	86.875781 0.1826	84.99922 0.5949
	4	Acc. time GI	99.898437 1.330626 1.130626	99.207813 3.2445	99.86766 0.7271
	5	Acc. time GI	99.4976563 1.4255 1.1593782	82.293750 4.1051	82.57734 0.8474

Table 6. Results of different amount of data for Configuration 3.

Set size	Adj. G		GI algorithm	EM algorithm	K-means
150	2	Ac. time GI	66.666667 0.0765 3.836915	66.666667 0.0195	66.66667 0.0109
	3	Acc. time GI	99.98 0.079 3.371382	99.973333 0.0085	97.30667 0.0288
	4	Acc. time GI	99.98 0.0811 3.371382	87.42 0.0231	84.70667 0.0301
300	2	Ac. time GI	66.666667 0.17 4.026724	66.666667 0.0154	66.66667 0.0094
	3	Acc. time GI	99.33667 0.1667 3.35724	99.336667 0.0115	98.00333 0.0389
	4	Acc. time GI	99.33667 0.2154 3.35724	87.226667 0.0295	85.31333 0.0538
1200	2	Ac. time GI	66.66667 0.2857 3.87596	66.3325 0.0365	66.66667 0.0187
	3	Acc. time GI	99.9125 0.3832 3.376	99.078333 0.0267	97.65833 0.1359
	4	Acc. time GI	99.9125 0.4729 3.376	87.972500 0.6719	85.18583 0.1533

is diagonal, and non-independent bivariate normal distributions, that is, Gaussians whose covariance matrix is a positive definite non-diagonal matrix. This is in order to verify how efficient the models are when we use databases that might not meet the independence condition used by the GI algorithm.

Configuration 1. (See Fig. 3(a)) We have 3000 data items generated with *different proportions* of 2 *separate bivariate* Gaussians, with the following parameters: 2000

data items with

$$\mu_1 = (-7, -3), \quad \Sigma_1 = \begin{pmatrix} 2 & -1 \\ -1 & 1 \end{pmatrix}$$

and 1000 data items with

$$\mu_2 = (3, 10), \quad \Sigma_2 = \begin{pmatrix} 1 & 1 \\ 1 & 2 \end{pmatrix}.$$

In Table 8 the results for this configuration when we adjust 2, 3 and 4 Gaussians are shown.

We obtain the best accuracy with the GI algorithm for 2, 3 and 4 Gaussians, and the EM algorithm and

the K -means method for 2 Gaussians. Once again, we obtained satisfactory results when adjusting 2 or more Gaussians with the GI algorithm because with this method we have values $\phi_k = 0$ for the extra components considered in the adjustment of the mixture. Also, in this configuration, we consider a generated data set from non-independent Gaussian distribution, which means that it does not comply with the independence condition, and we have favourable results.

Configuration 2. (See Fig. 3(b)) We have 3000 data points generated with *different proportions* of 4 bivariate Gaussians *intersected* by pairs, with the following parameters: 1200 data points with

$$\mu_1 = (0, 16), \quad \Sigma_1 = \begin{pmatrix} 2 & 0 \\ 0 & 1 \end{pmatrix},$$

900 data points with

$$\mu_2 = (7, 18), \quad \Sigma_2 = \begin{pmatrix} 1 & 1 \\ 1 & 2 \end{pmatrix},$$

600 data points with

$$\mu_3 = (8, -5), \quad \Sigma_3 = \begin{pmatrix} 2 & 0 \\ 0 & 2 \end{pmatrix}$$

and 300 data points with

$$\mu_4 = (14, -5), \quad \Sigma_4 = \begin{pmatrix} 1 & 0 \\ 0 & 2 \end{pmatrix}.$$

The results for this configuration, when we adjust 3, 4 and 5 Gaussians, are shown in Table 9.

Again, in this configuration we obtain the highest percentage for each model when we consider 4 Gaussians, and the GI algorithm has the highest average.

Configuration 3. (See Fig. 3(c)) We have 3000 data items generated with *equal proportions* from 4 *intersected bivariate* Gaussians, with the following parameters: 750 data items with

$$\mu_1 = (0, 16), \quad \Sigma_1 = \begin{pmatrix} 1 & -1 \\ -1 & 2 \end{pmatrix},$$

750 data items with

$$\mu_2 = (5, 16), \quad \Sigma_2 = \begin{pmatrix} 1 & 1 \\ 1 & 2 \end{pmatrix},$$

750 data items with

$$\mu_3 = (0, 10), \quad \Sigma_3 = \begin{pmatrix} 1 & 0 \\ 0 & 2 \end{pmatrix}$$

and 750 data items with

$$\mu_4 = (5, 11), \quad \Sigma_4 = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}.$$

Table 7. Results of different number of intervals in the representative histogram.

No. of intervals	Adj. G	Acc.	GI algorithm	EM algorithm	K -means
4	3	Acc. time GI	87.5 0.4551 5.710026	86.73533 0.1198	86.799 0.2823
	4	Acc. time GI	99.899667 0.5139 4.710025	94.907333 1.8622	99.832 0.365
	5	Acc. time GI	99.899667 0.9875 4.710025	83.800333 2.3186	81.07567 0.4648
20	3	Acc. time GI	87.5 0.4884 0.9650311	86.733333 0.1288	86.8 0.3123
	4	Acc. time GI	99.9 0.5834 0.3244576	94.966667 1.9476	99.83333 0.4083
	5	Acc. time GI	99.9 1.0465 0.3244576	83.766667 2.3792	81.06667 0.501
40	3	Acc. time GI	87.5 0.4737 0.8450669	86.73333 0.1386	86.8 0.2887
	4	Acc. time GI	99.9 0.5339 0.2116182	94.96667 1.89810	99.83333 0.3835
	5	Acc. time GI	99.9 0.9864 0.2116182	83.76667 2.3081	81.06667 0.468

Table 8. Results of Configuration 1: bivariate case.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	100 0.9672	100 0.0228	100 0.0105
3	Acc. time	100 1.5312	83.90367 0.1307	67.99533 0.0064
4	Acc. time	100 1.6254	52.33167 0.2826	62.93767 0.0072

Table 9. Results of Configuration 2: bivariate case.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
3	Acc. time	89.8144 1.0006	89.6008 0.1135	89.3394 0.0165
4	Acc. time	98.0784 1.0633	94.8428 0.7247	90.0292 0.0166
5	Acc. time	97.5006 1.3915	79.3204 2.8278	77.0836 0.0197

In Table 10 the obtained results are displayed.

In this case, we obtain high averages for 4 and 5 adjusted Gaussians with the GI algorithm. We obtain the second best result when we consider 4 Gaussians with the K -means method. With this example we can see that

Table 10. Results of Configuration 3: bivariate case.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
3	Acc. time	73.3720 1.9184	73.5376 0.8386	73.7796 0.0518
4	Acc. time	97.6964 1.749	93.4616 1.5722	97.1192 0.058
5	Acc. time	97.6964 1.3066	82.5796 4.7118	87.0752 0.0648

Table 11. Results of different amount of data for Configuration 2: bivariate case.

Set size	Adj. G		GI algorithm	EM algorithm	K -means
150	3	Ac. Time	89.89333 0.2482	89.34 0.0091	89.34 0.0074
	4	Acc. Time	99.99333 0.1834	99.34 0.0109	99.34 0.0055
	5	Acc. Time	99.98667 0.2792	85.3 0.0755	85.20667 0.0051
300	3	Ac. Time	88.67667 0.2811	89.66667 0.0141	89.66667 0.0061
	4	Acc. Time	99.99667 0.268	80.00333 0.0428	99.43 0.0059
	5	Acc. Time	99.99667 0.3704	69.9700 0.0497	82.76 0.0058
1200	3	Ac. Time	89.66833 0.7935	89.75167 0.0242	69.915 0.0049
	4	Acc. Time	99.91417 0.8211	99.75083 0.0444	99.58083 0.0062
	5	Acc. Time	99.91417 1.1526	87.90667 0.1197	85.9925 0.0062

the method is robust with respect to the independence condition established for the analyzed data set.

In the same way as in the univariate case, we will show some experiments considering a different amount of data. In this case, we think that the experiments with different numbers of intervals are not relevant since we got similar results to those observed in Table 7. The number of intervals does not affect the results. Furthermore, we can see that the algorithm performs well even though the independence hypothesis is not fulfilled.

Different amounts of data. For the experiments with different amounts of data, we use the same parameters for the means and covariance matrices of Configurations 2 and 3 in the bivariate case. We show the different amounts of data and the results in Table 11 for Configuration 2 and in Table 12 for Configuration 3. The proportions used correspond to the configuration ones, as in the univariate case.

Table 12. Results of different amount of data for Configuration 3: bivariate case.

Set size	Adj. G		GI algorithm	EM algorithm	K -means
160	3	Ac. time	73.2125 0.3401	74.8875 0.0108	74.9625 0.0061
	4	Acc. time	99.33875 0.2771	99.28125 0.0101	99.33125 0.0055
	5	Acc. time	99.33875 0.3963	83.0125 0.0137	88.70625 0.0053
320	3	Ac. time	68.73438 0.4153	73.44688 0.0102	73.44375 0.0041
	4	Acc. time	95.96563 0.3818	95.89062 0.0193	95.65625 0.003
	5	Acc. time	95.96563 0.6307	83.8250 0.0409	86.26562 0.0043
1280	3	Ac. time	62.73438 1.273	73.98438 0.0581	73.82812 0.0088
	4	Acc. time	97.89062 0.9137	96.07812 0.0941	97.8125 0.0073
	5	Acc. time	97.89062 1.5352	89.21875 0.1195	86.79688 0.0084

Table 13. Results for 2 Gaussians in dimension 5.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	100 5.6146	100 0.0969	100 0.0393
3	Acc. time	99.85305 6.6641	76.25725 8.4049	75.833 0.0642
4	Acc. time	97.23465 8.4456	67.4484 12.0507	66.45755 0.0843

4.1.3. Higher dimensions. In this section we consider experiments in dimensions 5 to 8 when we generate data from 2 and 3 Gaussians, in order to verify how efficient the models are in higher dimensions.

For this experiments we use the following values for the parameters: $\mu_1 = (1, 1, \dots, 1) \in \mathbb{R}^N$ and $\Sigma_1 = \text{Diag}(1, 1, \dots, 1)$ with size $N \times N$, $\mu_2 = (10, 10, \dots, 10) \in \mathbb{R}^N$ and $\Sigma_2 = \Sigma_1$. Remember that N is the corresponding dimension of the analyzed data set. For the cases with 3 Gaussians, $\mu_3 = (20, 20, \dots, 20) \in \mathbb{R}^N$ and $\Sigma_3 = \Sigma_1$.

In these cases we also transfer the data to

$$\underbrace{[-10, 10] \times \dots \times [-10, 10]}_{N \text{ times}},$$

taking into account the number of dimensions N .

Dimension 5. In Tables 13 and 14, we show the results obtained when we generated the data set with 2 and 3 Gaussians, respectively.

In the first case we obtained 100% of accuracy when we adjusted 2 Gaussians with the 3 models. We obtain the

Table 14. Results for 3 Gaussians in dimension 5.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	66.66667 5.508	66.66667 0.212	66.66667 0.074
3	Acc. time	99.98667 7.196	100 0.202	100 0.088
4	Acc. time	96.16133 8.92	84.44 21.052	83.442 0.118

Table 15. Results for 2 Gaussians in dimension 6.

Adjusted Gaussians		GI Algorithm	EM Algorithm	K -means
2	Acc. Time	100 0.7409	100 0.0115	100 0.0059
3	Acc. Time	100 1.0445	77.5225 0.9674	75.896 0.0054
4	Acc. Time	100 1.3269	70.5875 1.3206	68.2905 0.008

Table 16. Results for 3 Gaussians in dimension 6.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	66.66667 0.638	66.66667 0.033	66.66667 0.015
3	Acc. time	96.13667 0.833	93.68333 0.875	100 0.033
4	Acc. time	93.13333 1.086	60.35 1.216	83.86667 0.036

second best result when we adjust 3 Gaussians with the GI algorithm. In the second scenario, we have better results with the EM algorithm and the K -means method when we adjust 3 Gaussians. We obtain the second best result with the GI algorithm.

It is important to mention that if we do not have *a priori* information about the number of classes in the data set and we adjust a wrong number of classes, we obtain the best results with the GI algorithm.

Dimension 6. We show the results for the experiments in dimension 6 in Tables 15 and 16.

For the experiments with 2 Gaussians, we obtain better results when we adjust 2 Gaussians with 3 models and when we adjust 3 and 4 Gaussians with the GI algorithm. In this case we can observe that the GI algorithm guarantees better results even when we adjust more Gaussians because the model finds values equal to 0 for mixture proportions for the additional Gaussians. As for 3 Gaussians, we obtain better results with the K -means method.

As in the experiments in dimension 5, in these cases we obtain the best result if we adjust a wrong number of

Table 17. Results for 2 Gaussians in dimension 7.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	100 1.0257	100 0.012	100 0.0064
3	Acc. time	99.9975 1.3625	76.7625 0.9881	76.701 0.0066
4	Acc. time	99.9495 1.6538	72.1115 1.4013	68.392 0.0087

Table 18. Results for 3 Gaussians in dimension 7.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	66.66667 0.724	66.66667 0.024	66.66667 0.02
3	Acc. time	90.71333 0.93	90.70667 0.578	100 0.048
4	Acc. time	89.53333 1.212	87.377333 0.75	84.01333 0.028

Table 19. Results for 2 Gaussians in dimension 8.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	100 0.862	100 0.03	100 0.016
3	Acc. time	100 1.232	81.97 0.798	75.53 0.032
4	Acc. time	100 1.536	74.24 1.368	67.73 0.032

classes with the GI algorithm.

Dimension 7. In Table 17 we can observe the results obtained when we generated a data set from 2 Gaussians. The best results are obtained when we adjust 2 Gaussians with 3 models. In this case, we obtained some mixture proportions with values close to zero, so that the accuracy for 3 and 4 Gaussians is close to 100%.

For the case of 3 Gaussians, we obtain the best result with the K -means method, as can be seen in Table 18.

Dimension 8. In Table 19 we show the results for 2 Gaussians. Again, we obtain the best results when we adjust 2, 3 and 4 Gaussians with the GI algorithm and when we adjust 2 Gaussians with the 3 models.

For 3 Gaussians, we obtain the results in Table 20. The K -means method yields better results when we adjust 3 Gaussians. It is important to mention that with this model we only have values for the means of the data. Again, in these cases, if we adjust a wrong number of classes, we have the best result with the GI algorithm.

Remarks. Our chief remarks are as follows.

- In most cases we get better results with the GI algorithm than with the other two models, and when we do not get the best result, we get the second best.
- We obtain the best averages when considering the actual number of classes used to generate the data.
- We get better results when classes are well differentiated, graphically speaking.
- The difference between the average with the GI algorithm for the original number of classes and the following ones is very small, because, when we consider more classes, this model finds the parameters for the additional ones with proportion ϕ_k equal or close to 0.
- The Gini index values for the univariate case are lower when we consider the actual number of data classes.
- The model appears to be robust with respect to the number of data items and the number of intervals in the representative histogram.
- If we analyze the execution time of these methods, it is clear that the fastest and relatively effective method is the K -means ones; however, with it we cannot find values for the covariance matrices.
- When we compare the accuracy of the GI algorithm with that of the EM algorithm, the former performs better but takes longer to execute.
- We can ensure that the GI algorithm finds values of mixing proportions equal or close to zero when we fit more Gaussians than the actual amount used to generate the data.

4.2. Real data. To carry out classification experiments with real data, we consider the Iris data set and the Seeds data set of the UCI Machine Learning Repository. For each of these data sets we use the GI algorithm, the EM algorithm and the K -means method for data classification.

Similar results of these experiments are reported in our earlier work (López-Lobato and Avendaño-Garrido, 2020); however, here the execution times for the three considered models are added.

4.2.1. Iris data set. In the Iris data set the authors examine 3 different varieties of the Iris plant: Iris Setosa, Iris Versicolour and Iris Virginica, 50 instances each, considering 4 physical characteristics of this plants: sepal length, sepal width, petal length and petal width. In this case we have a 4-dimensional data set with 3 classes. The results obtained for this database are shown in Table 21.

We obtain the best percentage of accuracy for the 3 models when we consider 3 clusters, because the database has 3 differentiated classes. We obtain a better percentage of accuracy with the GI algorithm for 3 Gaussians.

Table 20. Results for 3 Gaussians in dimension 8.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	66.66667 0.8067	66.66667 0.0367	66.66667 0.0167
3	Acc. time	99.7 1.1267	58.38889 1.04333	100 0.0333
4	Acc. time	96.68889 1.4367	91.72222 1.47	83.97778 0.04

Table 21. Results for the Iris data set.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	66.66667 0.11	66.66667 2.39	66.66667 0.02
3	Acc. time	97 0.22	96.66667 3.36	88.66667 0.01
4	Acc. time	96 0.28	92 4.32	72.66667 0.02

Table 22. Results for the Seeds data set.

Adjusted Gaussians		GI algorithm	EM algorithm	K -means
2	Acc. time	66.66667 1.07	66.66667 2.68	66.19048 0.01
3	Acc. time	94.61905 1.27	93.33333 4.83	89.52381 0.02
4	Acc. time	92.85714 1.7	89.52381 9.97	77.14286 0.02

4.2.2. Seeds data set. In the Seeds data set, the authors examine 3 different varieties of wheat seeds: Kama, Rosa and Canadian, 70 instances each, considering 7 geometrical parameters of wheat grains: area, perimeter, compactness, length of kernel, width of kernel, asymmetry coefficient and length of kernel groove. Thus, we have a 7-dimensional data set with 3 classes.

By adjusting this data set through the GI algorithm, the EM algorithm, and the K -means method, we obtain the percentages of accuracy shown in Table 22.

With the three models we obtain a better percentage of accuracy when we consider three adjusted Gaussians. We obtain the best average with the GI algorithm.

It is important to mention that, in the experiments with real data, the time required by the GI algorithm is smaller than that required by the EM algorithm with better results. This leads us to the conclusion that the proposed model helps us to efficiently estimate the parameters of a Gaussian mixture model through the Gini index problem.

5. Conclusions and future work

Thanks to the experiments carried out in this work, we can say that with the proposed model we obtain favourable

results, because our model seeks to minimize the Gini index between the empirical distribution and the proposed parametric distribution.

With the proposed model we obtain good results, even if the independence condition is not met and if the analyzed data sets comes from distributions that are not Gaussian mixtures, as in Configurations 3 and 4 for the univariate case and the experiments with real data. Furthermore, if we do not know the number of classes present in the database and arbitrarily set this number, the model fits the true number of classes, as we observed with the simulated data.

As future work, we want to study the theoretical properties and convergence of the GI algorithm and search for applications with real data. Also, we want to find some way to automatically specify the number of classes present in the analyzed data set with our method, analogously to the articles by Kulczycki (2018) and Kłopotek *et al.* (2020).

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